

THE FUNDAMENTAL THEOREM OF MARKOV CHAINS

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ABSTRACT. This paper provides some background for and proves the Fundamental Theorem of Markov Chains. It provides some basic definitions and notation for recursion, periodicity, and stationary distributions.

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1. DEFINITIONS AND BACKGROUND

So what is a Markov Chain, let's define it.

Definition 1.1. Let $\{X_0, X_1, \dots\}$ be a sequence of random variables and $Z = 0, \pm 1, \pm 2, \dots$ be the union of the sets of their realizations. Then $\{X_0, X_1, \dots\}$ is called a *discrete-time Markov Chain* with state space Z if:

$$P(X_{n+1} = i_{n+1} | X_n = i_n, \dots, X_1 = i_1) = P(X_{n+1} = i_{n+1} | X_n = i_n)$$

Now let's set up some notation for the *one-step transition probabilities* of the Markov Chain Let:

$$p_{ij}(n) = P(X_{n+1} = j | X_n = i); \quad n = 0, 1, \dots$$

We will limit ourselves to *homogeneous* Markov Chains. Or Markov Chains that do not evolve in time.

Definition 1.2. We say that a Markov Chain is *homogeneous* if its one-step transition probabilities do not depend on n ie.

$$\forall n, m \in \mathbb{N} \quad \text{and} \quad i, j \in Z \quad p_{ij}(n) = p_{ij}(m)$$

We then define the *n-step transition probabilities* of a homogeneous Markov Chain by

$$p_{ij}^{(m)} = P(X_{n+m} = j | X_n = i)$$

Where by convention we define

$$p_{ij}^{(0)} = \begin{cases} 1 & \text{if } i = j \\ 0 & \text{if } i \neq j \end{cases}$$

With these definitions we are now ready for a quick theorem: the *Chapman-Kolmogorov equation*. This serves to relate the transition probabilities of a Markov Chain.

Theorem 1.3.

$$p_{ij}^{(m)} = \sum_{k \in \mathbf{Z}} p_{ik}^{(r)} p_{kj}^{(m-r)}; \quad \forall r \in N \cup \{0\}$$

Proof. Making use of the total probability rule as well as the Markovian property in this proof yields:

By total probability rule:

$$\begin{aligned} p_{ij}^{(m)} &= P(X_m = j | X_0 = i) = \sum_{k \in \mathbf{Z}} P(X_m = j, X_r = k | X_0 = i) \\ &= \sum_{k \in \mathbf{Z}} P(X_m = j | X_r = k, X_0 = i) P(X_r = k | X_0 = i) \end{aligned}$$

and now by the Markovian property:

$$= \sum_{k \in \mathbf{Z}} P(X_m = j | X_r = k) P(X_r = k | X_0 = i) = \sum_{k \in \mathbf{Z}} p_{ik}^{(r)} p_{kj}^{(m-r)}$$

□

The Chapman-Kolmogorov equation allows us to associate our n-step transition probabilities to a matrix of n-step transition probabilities. Lets define the matrix:

$$P^{(m)} = ((p_{ij}^{(m)}))$$

Corollary 1.4. Observe now that:

$$P^{(m)} = P^m$$

Where P^m here denotes the typical matrix multiplication.

Proof. We can rewrite the Chapman-Kolmogorov equation in matrix form as:

$$P^{(m)} = P^{(r)} P^{(m-r)}; \quad \forall r \in N \cup \{0\}$$

Now we can induct on m. It's apparent that

$$P^{(1)} = P$$

Giving us:

$$P^{(m+1)} = P^{(r)} P^{(m+1-r)}$$

□

A Markov Chain's transition probabilities along with an initial distribution completely determine the chain.

Definition 1.5. An initial distribution is a probability distribution $\{\pi_i = P(X_0 = i) | i \in \mathbf{Z}\}$ $\sum_{i \in \mathbf{Z}} \pi_i = 1$

Such a distribution is said to be stationary if it satisfies

$$\pi_j = \sum_{i \in \mathbf{Z}} \pi_i p_{ij}$$

Definition 1.6. We say that a subset space state $\mathbf{C} \subset \mathbf{Z}$ is *closed* if

$$\sum_{j \in \mathbf{C}} p_{ij} = 1 \quad \forall i \in \mathbf{C}$$

If \mathbf{Z} itself has no proper closed subsets then the Markov Chain is said to be *irreducible*.

Definition 1.7. We define the *period* of a state i by

$$d_i = \gcd(m \in \mathbb{Z} | p_{ii}^{(m)} > 0)$$

State i is *aperiodic* if $d_i = 1$

Definition 1.8. A state i is said to be *accessible* from a state j if there is an $m \geq 1$ s.t

$$p_{ij}^{(m)} > 0$$

If i is accessible from j and j is accessible from i then i and j are said to *communicate*.

Definition 1.9. We define the *first-passage time probabilities* by

$$f_{ij}^{(m)} = P(X_m = j; X_k \neq j, 0 < k < m - 1 | X_0 = i); \quad i, j \in \mathbb{Z}$$

And we will denote the expected value, the expected return time, of this distribution by

$$\mu_{ij} = \sum_{m=1}^{\infty} m f_{ij}^{(m)}$$

Definition 1.10. We say that a state i is *recurrent* if

$$\sum_{m=1}^{\infty} f_{ij}^{(m)} = 1$$

and *transient* if

$$\sum_{m=1}^{\infty} f_{ij}^{(m)} < 1$$

A recurrent state i is *positive-recurrent* if $\mu_{ii} < \infty$ and *null-recurrent* if $\mu_{ii} = \infty$

2. FUNDAMENTAL THEOREM OF MARKOV CHAINS

Theorem 2.1. For any irreducible, aperiodic, positive-recurrent Markov Chain there exist a unique stationary distribution $\{\pi_j, j \in \mathbb{Z}\}$ s.t $\forall i \in \mathbb{Z}$

Proof. Because our chain is irreducible, aperiodic, and positive-recurrent we know that for all $i \in \mathbb{Z}$ $\pi_j = \lim_{n \rightarrow \infty} p_{ij}^{(n)} > 0$. Likewise as π_j is a probability distribution

$\sum_{j \in \mathbb{Z}} \pi_j = 1$. We know that for any m

$$\sum_{i=0}^m p_{ij}^{(m)} \leq \sum_{i=0}^{\infty} p_{ij}^{(m)} \leq 1$$

taking the limit

$$\lim_{m \rightarrow \infty} \sum_{i=0}^m p_{ij}^{(m)} = \sum_{i=0}^{\infty} \pi_j \leq 1$$

which implies that for any M

$$\sum_{i=0}^M \pi_j \leq 1$$

Now we can use the *Chapman-Kolmogorov* equation analogously

$$p_{ij}^{(m+1)} = \sum_{i=0}^{\infty} p_{ik}^{(m)} p_{kj} \geq \sum_{i=0}^M p_{ik}^{(m)} p_{kj}$$

as before we take the limit as $m, M \rightarrow \infty$ yielding

$$\pi_j \geq \sum_{i=0}^{\infty} \pi_k p_{kj}$$

Now we search for contradiction and assume strict inequality holds for at least one state j . Summing over these inequalities yields

$$\sum_{j=0}^{\infty} \pi_j > \sum_{j=0}^{\infty} \sum_{k=0}^{\infty} \pi_k p_{kj} = \sum_{k=0}^{\infty} \pi_k \sum_{j=0}^{\infty} p_{kj} = \sum_{k=0}^{\infty} \pi_k$$

But this is a contradiction, hence equality must hold

$$\pi_j = \sum_{k=0}^{\infty} \pi_k p_{kj}$$

Thus a unique stationary distribution exists. \square

Futher it can be shown that this unique distribution is related to the expected return value of the Markov chain by

$$\pi_j = \frac{1}{\mu_{jj}}$$

REFERENCES

- [1] F. E. Beichelt. L. P. Fatti. Stochastic Processes And Their Applications. Taylor and Francis. 2002.